A Comparison of Student Skill Knowledge Estimates

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A fundamental goal of educational research is identifying students' current stage of skill mastery (complete/partial/none). In recent years a number of cognitive diagnosis models have become a popular means of estimating student skill knowledge. However, these models become difficult to estimate as the number of students, items, and skills grows. There exist alternatives such as sum-scores and the capability matrix. While initial theoretical work on sum-scores has been done by, the behavior of sum-scores and the capability matrix is not well understood with respect to each other or to estimates from cognitive diagnosis models. In this paper we compare the performance of the three estimates of student skill skill knowledge under a variety of clustering methods.

1 Introduction

A fundamental goal of educational research is identifying students' current stage of skill mastery (complete/partial/none). In addition, finding groups of students with similar skill set profiles is important to provide feedback for classroom instruction. In recent years a number of cognitive diagnosis models [3,9] have become a popular means of estimating student skill knowledge. However, these models become difficult and time-consuming to estimate as the number of students, items, and skills increases [6]. There are a variety of other procedures, such as sum-scores [3,7] and the capability matrix [1], that can be used to estimate student skill knowledge in (near to) real time.

While initial theoretical work on sum-scores has been done by [3], the behavior of sum-scores and the capability matrix is not well understood with respect to each other or to estimates from cognitive diagnosis models. In this paper we take step back and compare the performance of the three estimates of student skill skill knowledge under a variety of clustering methods.

In Section 2, we describe the three different estimates of student skill knowledge we focus on in this paper. In Section 3, we give a brief introduction to clustering methods. In Section 4, we show results from a simulation study. Finally, in Section 5, we offer conclusions and thoughts on future work.

2 Estimates of Student Skill Knowledge

When estimating student skill knowledge, there are many possible methods. This paper will consider three different estimation procedures. First, we introduce notation that will be common among the methods. We begin by assembling the skill dependencies of each item into a Q-matrix [2,13]. The Q-matrix, also referred to as a transfer model or skill

coding, is a $J \times K$ matrix where $q_{jk} = 1$ if item *j* requires skill *k* and 0 if it does not, *J* is the total number of items, and *K* is the total number of skills. The *Q*-matrix is usually an expert-elicited assignment matrix. This paper assumes the *Q*-matrix is known and correct.

There are (at least) two ways in which *Q*-matrices can differ. First, each item could require only a single skill or multiple skills. A Q-matrix can then be comprised of all single skill items, single and multiple skill items, or all multiple skill items. Second, the Q-matrix may have a balanced or unbalanced design. In a balanced design, all single skill items occur the same number of times and each combination of skills occurs the same number of times. For example, if K = 3 and J = 30 one possible balanced design would be five single skill items for each skill, four double skill items for each pair of skills, and three triple skill items. A design could be unbalanced in two ways. Either all skills or combinations of skills are present but do not occur the same number of times or there are missing skills or combinations of skills.

$$Q = \begin{bmatrix} q_{1,1} & q_{1,2} & \dots & q_{1,K} \\ \vdots & \ddots & & \vdots \\ q_{J,1} & q_{J,2} & \dots & q_{J,K} \end{bmatrix}, \quad Y = \begin{bmatrix} y_{1,1} & y_{1,2} & \dots & y_{1,J} \\ \vdots & \ddots & & \vdots \\ y_{N,1} & y_{N,2} & \dots & y_{N,J} \end{bmatrix}$$

We then assemble student responses in a $N \times J$ response matrix Y where y_{ij} indicates both if student i attempted item j and whether or not they answered item j correctly and N is the total number of students. If student i did not answer item j then $y_{ij} = NA$. The indicator $I_{y_{ij}\neq NA} = 0$ expresses this missing value. If student i attempted item j ($I_{y_{ij}\neq NA} = 1$), then $y_{ij} = 1$ if they answered correctly, or 0 if they answered incorrectly.

2.1 DINA Model Estimates

The first method of estimating student skill knowledge uses a common conjunctive cognitive diagnosis model. The deterministic inputs, noisy "and" gate model (DINA; [9]) models student responses as

$$P(Y_{ij} = 1 \mid \eta_{ij}, s_j, g_j) = (1 - s_j)^{\eta_{ij}} g_j^{1 - \eta_{ij}}$$
(1)

where $\alpha_{ik} = I_{\{\text{Student } i \text{ has skill } k\}}$ indicates if student *i* possesses skill *k*, $\eta_{ij} = \prod_{k=1}^{K} \alpha_{ik}^{q_{jk}}$ indicates if student *i* has all skills needed for item *j*, for item *j*, $s_j = P(Y_{ij} = 0 | \eta_{ij} = 1)$ is the slip parameter and $g_j = P(Y_{ij} = 1 | \eta_{ij} = 0)$ is the guess parameter. If a student is missing any of the required skills, the probability that they will answer an item correctly drops due to the conjunctive assumption.

We estimate the student skill knowledge parameters of the DINA model, the α_{ik} , using Markov Chain Monte Carlo methods with the program WinBUGS (Bayesian Inference Using Gibbs Sampling, [10]). In the model, the α_{ik} are 0/1 indicating whether or not student *i* has mastered skill *k*. Our estimates will be $\hat{\alpha}_{ik} \in [0, 1]$. We can think of the $\hat{\alpha}_{ik}$ as the probability that student *i* has mastered skill *k*.

2.2 Sum-scores

The second method we consider is the sum-score method of [4,7]. Here the variable $W_i = (W_{i1}, W_{i2}, ..., W_{iK})$ is defined as a vector of sum-scores where the k^{th} component is defined as

$$W_{ik} = \sum_{j=1}^{J} Y_{ij} q_{jk}.$$
 (2)

Thus, the components of W_i are simply the number of items student *i* answered correctly for each skill *k*. When an item requires more than one skill it will contribute to more than one component of W_i .

2.3 Capability Matrix

Finally, we consider the *capability matrix* we defined in [1]. The capability matrix B is an $N \times K$ matrix where where B_{ik} is the proportion of correctly answered items involving skill k that student i attempted. Thus,

$$B_{ik} = \frac{\sum_{j=1}^{J} I_{y_{ij} \neq NA} \cdot y_{ij} \cdot q_{jk}}{\sum_{j=1}^{J} I_{y_{ij} \neq NA} \cdot q_{jk}}$$
(3)

where y_{ij} and q_{jk} are the corresponding entries from the response matrix Y and Q-matrix. The capability matrix expands on the sum-score method by accounting for the number of items requiring skill k that student i answered. In this manner the statistic scales for the number of items in which the skill appears as well as for missing data. If a student has not seen all of the items requiring a particular skill, we still derive an estimate based on the available information. If student i completes no items involving skill k, then $B_{ik} = NA$. In this case, we impute an uninformative value (e.g., 0.5, mean, median) to map students to the hypercube. Exploring the performance of these imputation choices is ongoing. For this paper we assume that the data are complete or that missing B-values are appropriately imputed.

We can note that both the DINA model estimates and the *B*-matrix values map students into a *K*-dimensional hypercube (for each dimension, zero indicates total lack of skill mastery, one is complete skill mastery, and values in between are less certain). The 2^{K} corners of the hypercube correspond to natural skill set profiles $C_i = \{C_{i1}, C_{i2}, ..., C_{iK}\}, C_{ik} \in \{0, 1\}$.

Additionally, we can note theoretical connections between the sum-scores and *B*-matrix values. When all students have answered all questions and there is a balanced *Q*-matrix design, the two estimates will lie in the same feature space. In this case, we expect the two estimates to perform similarly. However, when there is either missing data or an unbalanced *Q*-matrix design, the space in which the estimates lie will be different. In this case, we can not guarantee that performance will be similar.

3 Clustering Methods

To identify groups of students with similar skill set profiles, we cluster the student skill knowledge estimates. In this paper we will compare the performance of three common clustering methods: hierarchical agglomerative clustering, K-means, and model-based clustering. In the sections below we will briefly introduce each of these methods.

3.1 Hierarchical Agglomerative Clustering

Hierarchical agglomerative clustering (HAC; [11]) links groups in order of closeness to form a tree structure from which a clustering solution can be extracted. Euclidean distance is most commonly used to measure the distance between groups. The method also requires the user to specify how to measure the distance between groups. We will use "complete" linkage where the distance between any two groups is defined as the largest distance between observations in the two groups. In HAC, all points begin as their own observation. The two closest groups are merged and all inter-group distances are recalculated. We continue merging groups and recalculating distances until a single group with all observations is formed. One the tree structure is formed, we can extract the desired number of clusters G by cutting the tree at a height corresponding to G branches.

3.2 K-means

k-Means [5] is a popular iterative descent algorithm for data $X = \{\underline{x}_1, \underline{x}_2, ..., \underline{x}_n\}, \underline{x}_i \in \Re^K$. It uses squared Euclidean distance as a dissimilarity measure and tries to minimize withincluster distance and maximize between-cluster distance. For a given number of clusters G, k-Means searches for cluster centers m_g and assignments A that minimize the criterion

$$\min_{A} \sum_{g=1}^{G} \sum_{A(i)=g} ||\underline{x}_i - m_g||^2.$$

The algorithm alternates between optimizing the cluster centers for the current assignment (by the current cluster means) and optimizing the cluster assignment for a given set of cluster centers (by assigning to the closest current center) until convergence (i.e. cluster assignments do not change). It tends to find compact, spherical clusters and requires *a priori* both the number of clusters *G* and a starting set of cluster centers. The final cluster assignment can be sensitive to the choice of centers; a common method for initializing k-Means is to randomly choose *G* observations.

3.3 Model-based Clustering

Model-based clustering [4, 12] is a parametric statistical approach that assumes: the data $X = \{\underline{x}_1, \underline{x}_2, ..., \underline{x}_n\}, \underline{x}_i \in \mathbb{R}^K$ are an independently and identically distributed sample from an unknown population density $p(\underline{x})$; each population group g is represented by a (often Gaussian) density $p_g(\underline{x})$; and $p(\underline{x})$; each population density components, i.e. $p(\underline{x}) = \sum_{g=1}^G \pi_g \cdot p_g(\underline{x}; \theta_g)$ where $\sum \pi_g = 1, 0 < \pi_g \leq 1$ for g = 1, 2, ..., G, and $\theta_g = (\mu_g, \Sigma_g)$ for Gaussian components. The method chooses the number of components G by maximizing the Bayesian Information Criterion (BIC) and estimates the means and variances (μ_g, Σ_g) via maximum likelihood. While it may assume Gaussian components, its

| Ν | J | Κ | Q-matrix design | DINA | HAC | K-means | MBC | MBC 2^{K} |
|-----|----|---|-------------------|----------|----------|----------|----------|-------------|
| 250 | 30 | 3 | Both, bal | 0.9793 | 0.9781 | 0.8367 | 0.8915 | 0.9632 |
| | | | | (0.0179) | (0.0200) | (0.1192) | (0.0882) | (0.1087) |
| 250 | 30 | 3 | Both,unbal, all | 0.9657 | 0.9657 | 0.7789 | 0.9129 | 0.9350 |
| | | | | (0.0285) | (0.2920) | (0.0941) | (0.0505) | (0.0758) |
| 250 | 30 | 3 | Both, unbal, miss | 0.9240 | 0.9131 | 0.7696 | 0.8811 | 0.9132 |
| | | | | (0.0395) | (0.0427) | (0.0858) | (0.0696) | (0.0428) |
| 250 | 30 | 3 | Mult, bal | 0.4677 | 0.5127 | 0.5012 | 0.5282 | 0.4979 |
| | | | | (0.0292) | (0.0443) | (0.0578) | (0.0690) | (0.0411) |
| 250 | 30 | 3 | Mult, unbal, all | 0.4629 | 0.4874 | 0.4948 | 0.5130 | 0.4790 |
| | | | | (0.0430) | (0.0536) | (0.0816) | (0.0736) | (0.0495) |
| 250 | 30 | 3 | Mult, unbal, miss | 0.3239 | 0.4070 | 0.3835 | 0.4266 | 0.4090 |
| | | | | (0.0380) | (0.0596) | (0.0521) | (0.0837) | (0.0630) |
| 500 | 68 | 5 | Both, bal | 0.9463 | 0.9428 | 0.7132 | 0.8348 | 0.9243 |
| | | | | (0.0184) | (0.0188) | (0.0428) | (0.1123) | (0.0488) |
| 500 | 68 | 5 | Both, unbal, miss | 0.8724 | 0.8729 | 0.6665 | 0.8213 | 0.8624 |
| | | | | (0.0247) | (0.0219) | (0.0466) | (0.0960) | (0.0226) |
| 300 | 40 | 7 | Single | 0.9041 | 0.8891 | 0.7674 | 0.3050 | 0.8881 |
| | | | | (0.0262) | (0.0286) | (0.0409) | (0.1203) | (0.0282) |

Table 1: Clustering Results Based on DINA Model Estimates of Student Skill Knowledge

flexibility on their shape, volume, and orientation allows student groups of varying shapes and sizes. When multiple students may map to the same location model-based clustering is known to overfit the data by using spikes with near singular covariance in these locations. To alleviate this concern, we jitter the student skill estimates by a small amount (0.01). The effect on our results is minimal.

4 Simulation Study

To compare the skill knowledge estimates and clustering methods described above we did a simulation study, generating data from the DINA model, described in Equation 1. We first fix skill difficulties to be of equal medium difficulty and set inter-skill correlation to be zero and generate true skill set profiles C_i for each student. These parameter choices evenly spread students among the 2^K natural skill set profiles. Next we draw slip and guess parameters from a random uniform distribution ($s_j \sim \text{Unif}(0,0.30)$; $g_j \sim \text{Unif}(0,0.15)$). Given profiles and slip/guess parameters, we generate the student response matrix Y. We repeat the data generation using a variety of different Q-matrices. For each Q-matrix we ran 20 simulations so that we could derive an estimate of the standard deviation.

For these examples we know the true underlying skill set profiles C_i and can calculate their agreement with the clustering partitions using the Adjusted Rand Index (ARI; [8]), a common measure of agreement between two partitions. The expected value of the ARI is zero and the maximum value is one, with larger values indicating better agreement.

Tables 1, 2, and 3 show the clustering results for the DINA model estimates, sum-scores,

| Ν | J | Κ | Q-matrix design | HAC | K-means | MBC | MBC 2^{K} |
|-----|----|---|-------------------|----------|----------|----------|-------------|
| 250 | 30 | 3 | Both, bal | 0.7644 | 0.8156 | 0.9321 | 0.9442 |
| | | | | (0.1095) | (0.1110) | (0.1181) | (0.0515) |
| 250 | 30 | 3 | Both, unbal, all | 0.6398 | 0.7707 | 0.6970 | 0.8494 |
| | | | | (0.0889) | (0.0951) | (0.2138) | (0.0713) |
| 250 | 30 | 3 | Both,unbal,miss | 0.6482 | 0.6728 | 0.7066 | 0.7661 |
| | | | | (0.0511) | (0.0650) | (0.2064) | (0.1095) |
| 250 | 30 | 3 | Mult, bal | 0.3950 | 0.4720 | 0.4383 | 0.4375 |
| | | | | (0.0339) | (0.0648) | (0.0675) | (0.0517) |
| 250 | 30 | 3 | Mult, unbal, all | 0.3862 | 0.4606 | 0.4380 | 0.4481 |
| | | | | (0.0533) | (0.0670) | (0.0696) | (0.0428) |
| 250 | 30 | 3 | Mult, unbal, miss | 0.2689 | 0.2827 | 0.3314 | 0.3099 |
| | | | | (0.0273) | (0.0848) | (0.0352) | (0.0347) |
| 500 | 68 | 5 | Both, bal | 0.4006 | 0.5859 | 0.5893 | 0.6523 |
| | | | | (0.0560) | (0.0442) | (0.1223) | (0.0432) |
| 500 | 68 | 5 | Both, unbal, miss | 0.4104 | 0.54412 | 0.6010 | 0.6265 |
| | | | | (0.0373) | (0.0366) | (0.0537) | (0.0397) |
| 300 | 40 | 7 | Single | 0.7348 | 0.6474 | (0.0973) | 0.7080 |
| | | | | (0.0526) | (0.0456) | (0.0362) | (0.0453) |

Table 2: Clustering Results Based on Sum-scores Estimates of Student Skill Knowledge

and the capability matrix, respectively. In each table, *N* is the number of students, *J* is the number of items, and *K* is the number of skills. In the DINA column of Table 1, we rounded the $\hat{\alpha}_{ik}$ to 0/1, found the closest skill set profile, and compared it to the true generating skill set profile. In the remaining columns we clustered the unrounded $\hat{\alpha}_{ik}$. In Tables 2 and 3 we cluster the sum-score vectors and capability matrix. The final two columns in each of the tables gives the results for model-based clustering when we search over an appropriate range and when we request 2^{K} clusters.

When looking at the tables we can note that the performance of all three clustering methods is better (as indicated by a higher ARI) when there are both single and multiple skill items in the *Q*-matrix. In addition, when the *Q*-matrix has a balanced design, as opposed to an unbalanced design, the recovery of the true skill set profiles is better. Finally, we can note that the performance of the three estimates of the student skill set profiles is similar across the clustering methods. It is interesting to note that they perform as well as the DINA model estimates.

5 Conclusions

Simulated examples show that for Q-matrices with single and multiple skill items, the recovery of the true skill set profiles was better for all three clustering methods and all three methods of estimating student skill knowledge. In addition, we note that the alternative methods of estimating student skill knowledge behave similarly. It is interesting to note that they perform as well as the DINA model estimates.

| Ν | J | Κ | Q-matrix design | HAC | K-means | MBC | MBC 2^{K} |
|-----|----|---|-------------------|----------|----------|----------|-------------|
| 250 | 30 | 3 | Both, bal | 0.7644 | 0.7947 | 0.9353 | 0.9411 |
| | | | | (0.1095) | (0.1056) | (0.1583) | (0.0300) |
| 250 | 30 | 3 | Both, unbal, all | 0.7273 | 0.8082 | 0.6252 | 0.8281 |
| | | | | (0.0867) | (0.1227) | (0.1719) | (0.1543) |
| 250 | 30 | 3 | Both,unbal,miss | 0.6698 | 0.7390 | 0.4563 | 0.6693 |
| | | | | (0.0813) | (0.0778) | (0.1267) | (0.1628) |
| 250 | 30 | 3 | Mult, bal | 0.4045 | 0.4530 | 0.4586 | 0.4499 |
| | | | | (0.0347) | (0.0508) | (0.0624) | (0.0382) |
| 250 | 30 | 3 | Mult, unbal, all | 0.3899 | 0.4585 | 0.4518 | 0.4580 |
| | | | | (0.0509) | (0.0550) | (0.0822) | (0.0589) |
| 250 | 30 | 3 | Mult, unbal, miss | 0.2700 | 0.3638 | 0.2803 | 0.2840 |
| | | | | (0.0291) | (0.0737) | (0.0620) | (0.0457) |
| 500 | 68 | 5 | Both, bal | 0.4096 | 0.5711 | 0.5951 | 0.6647 |
| | | | | (0.0504) | (0.0543) | (0.1284) | (0.0928) |
| 500 | 68 | 5 | Both, unbal, miss | 0.4327 | 0.5435 | 0.5560 | 0.6291 |
| | | | | (0.0405) | (0.0350) | (0.2027) | (0.1050) |
| 300 | 40 | 7 | Single | 0.7399 | 0.6437 | 0.0906 | 0.7109 |
| | | | | (0.0545) | (0.0402) | (0.0168) | (0.0409) |

Table 3: Clustering Results Based on Capability Matrix Estimates of Student Skill Knowledge

While there are benefits of using the capability matrix, we can note that if an item requires multiple skills and a student answers incorrectly, all skills required by the item will receive a penalty, even if the student has mastered one (or more) of the skills. In future work we will explore the behavior of alternative capability matrices that better account for multiple skill items. Possible methods could use performance on single skill items or simply weight by the number of skills required by the incorrectly answered item.

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